

# Omega-3 polyunsaturated fatty acid supplementation and cognition: A systematic review and meta-analysis

Ruth E Cooper, Charlotte Tye, Jonna Kuntsi, Evangelos Vassos\* and Philip Asherson\*

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## Abstract

**Background:** Omega-3 polyunsaturated fatty acids (*n*-3 PUFAs) are promoted as cognitive enhancers with consumption recommended in the general population and those with neurocognitive deficits such as attention deficit hyperactivity disorder (ADHD). However, evidence from randomised placebo-controlled trials is inconclusive.

**Aims:** This study aimed to conduct a systematic review and meta-analysis examining the effect of *n*-3 PUFA supplementation on cognition in healthy populations and those with ADHD and related disorders (RDs).

**Methods:** Databases were searched for randomised controlled trials (RCTs) in adults and school-aged children (who were healthy and typically developing (TD) or had ADHD or a related-neurodevelopmental disorder (ADHD+RD) which assessed the effects of *n*-3 PUFA on cognition.

**Results:** In the 24 included studies *n*-3 PUFA supplementation, in the whole sample and the TD and ADHD+RD subgroup, did not show improvements in any of the cognitive performance measures. In those with low *n*-3 PUFA status, supplementation improved short-term memory.

**Conclusions:** There is marginal evidence that *n*-3 PUFA supplementation effects cognition in those who are *n*-3 PUFA deficient. However, there is no evidence of an effect in the general population or those with neurodevelopmental disorders. This has important implications given the widespread advertisement and consumption of *n*-3 PUFA; claims of cognitive benefit should be narrowed.

## Keywords

Attention deficit hyperactivity disorder, cognition, omega-3, randomised controlled trial, meta-analysis

## Introduction

Global spending on omega-3 products is in the billions with consumption recommended in both the general population and those with neurocognitive deficits such as attention deficit hyperactivity disorder (ADHD; Bloch and Qawasmi, 2011), psychosis (Amming et al., 2010), depression (Su et al., 2014) and autism (Yui et al., 2012). Stimulant medications significantly reduce the symptoms and cognitive impairments in ADHD (Banaschewski et al., 2006; Coghill et al., 2014; Faraone and Buitelaar, 2010). However some individuals elect against such medication due to undesirable side-effects, partial response and questions regarding the long-term efficacy and developmental effects (Dunnick and Hailey, 1995; Leonard et al., 2004; Nasrallah et al., 1986). Omega-3 polyunsaturated fatty acid (*n*-3 PUFA) supplementation is an extensively studied alternative treatment for ADHD, with meta-analyses of behavioural data demonstrating a small but significant effect on ADHD symptom improvement in children (Bloch and Qawasmi, 2011; Sonuga-Barke et al., 2013). It has also been proposed that *n*-3 PUFA supplements are important for the health of the brain and improve cognitive functions (Bryan et al., 2004). However, as yet there has been no systematic evaluation of the available evidence on which to draw any firm conclusions about its efficacy.

Longitudinal and cross sectional studies suggest an association between increased *n*-3 PUFA intake and cognitive function (Aberg et al., 2009; Bryan et al., 2004; Hibbeln et al., 2007). One

of the main explanations proposed is based on the high lipid cell membrane composition, maintenance of which may be vital for the optimal development and function of the brain and nervous system (Bryan et al., 2004). However, randomised controlled trials (RCTs) in typically developing (TD) participants and those with ADHD and related neurodevelopmental disorders, have instead yielded mixed results. Benefits of *n*-3 PUFA supplementation on cognitive performance have been reported in healthy adults (Stonehouse et al., 2013) and children with ADHD (Sinn et al., 2008) and developmental coordination disorder (DCD) (Richardson and Montgomery, 2005). Yet a number of other studies in these populations have failed to find an effect (Jackson et al., 2012; Kairaluoma et al., 2009; Milte et al., 2012; Osendarp et al., 2007)

King's College London, MRC Social, Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology and Neuroscience, London London, UK

\*Co-senior authors.

## Corresponding author:

Ruth E Cooper, King's College London, MRC Social, Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology and Neuroscience, London SE5 8AF, UK.  
Email: ruth.cooper@kcl.ac.uk

Given the global market for omega-3 products it is of public importance that there is a more conclusive picture as to whether *n*-3 PUFA supplementation improves cognitive performance. We therefore conducted a systematic review and meta-analysis of randomised placebo-controlled trials which examined the effect of *n*-3 PUFA supplementation on cognitive performance in healthy populations and those with ADHD and related neurodevelopmental disorders.

## Methods

We conducted a systematic review and meta-analysis in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines (Moher et al., 2009).

### Eligibility criteria and data extraction

Studies were included if: (a) they were randomised double-blind placebo-controlled trials of *n*-3 PUFA supplementation including docosahexaenoic acid (DHA), eicosapentaenoic acid (EPA) and alpha-linolenic acid (ALA). Trials supplementing with ALA alone were excluded as ALA is thought to have a limited impact on cognition compared to EPA and DHA (Kalmijn et al., 2004) and humans have a limited capacity to synthesise EPA and DHA from ALA (Goyens et al., 2005); (b) participants were school-aged children (4–12 years), adolescents (13–17 years) or adults (18–60 years) who were either healthy (TD group) or had a diagnosis of ADHD or high levels of ADHD symptoms or related neurodevelopmental traits such as DCD or dyslexia (ADHD+RD group); and (c) the study measured cognitive performance defined as (one or more measure of): intelligence quotient (IQ), inhibition, attention (omission errors), working memory, short-term memory, reading, spelling, mean reaction time and reaction time variability (see Supplementary Material, Table S1 for details). There were no language restrictions on trial eligibility.

The databases Ovid Medline (1946–September, week 2, 2014), Embase (1974–2014, week 37) and Psychinfo (1806–September, week 3, 2014) were searched. References of eligible trials and appropriate reviews were searched for additional citations. Unpublished or ongoing trials were searched on the ClinicalTrials.gov website and authors contacted to request relevant data. The search was updated in November 2014. The search terms used are listed in Supplementary Material, Table S2.

Risk of bias to determine study quality was assessed independently by two authors (REC and CT) according to PRISMA guidelines and the Cochrane Handbook of Systematic Reviews (Higgins and Green, 2011) (Supplementary Material, Table S3 and S4). Decision to include was based on risk of bias which was classed as low, unclear or high. Unresolved classification of studies was arbitrated by PA.

Data extraction was performed by REC and checked by a research assistant. The main outcome measures were the mean and standard deviation (SD) of the pre and post treatment cognitive performance measures for active and placebo arms, with intent to treat (ITT) analysis preferentially reported. Additional measures investigated included participant characteristics, study design and the supplement type and dose. If multiple treatment arms were present, only those supplementing with *n*-3 PUFA or placebo were included. With regard to missing data, we

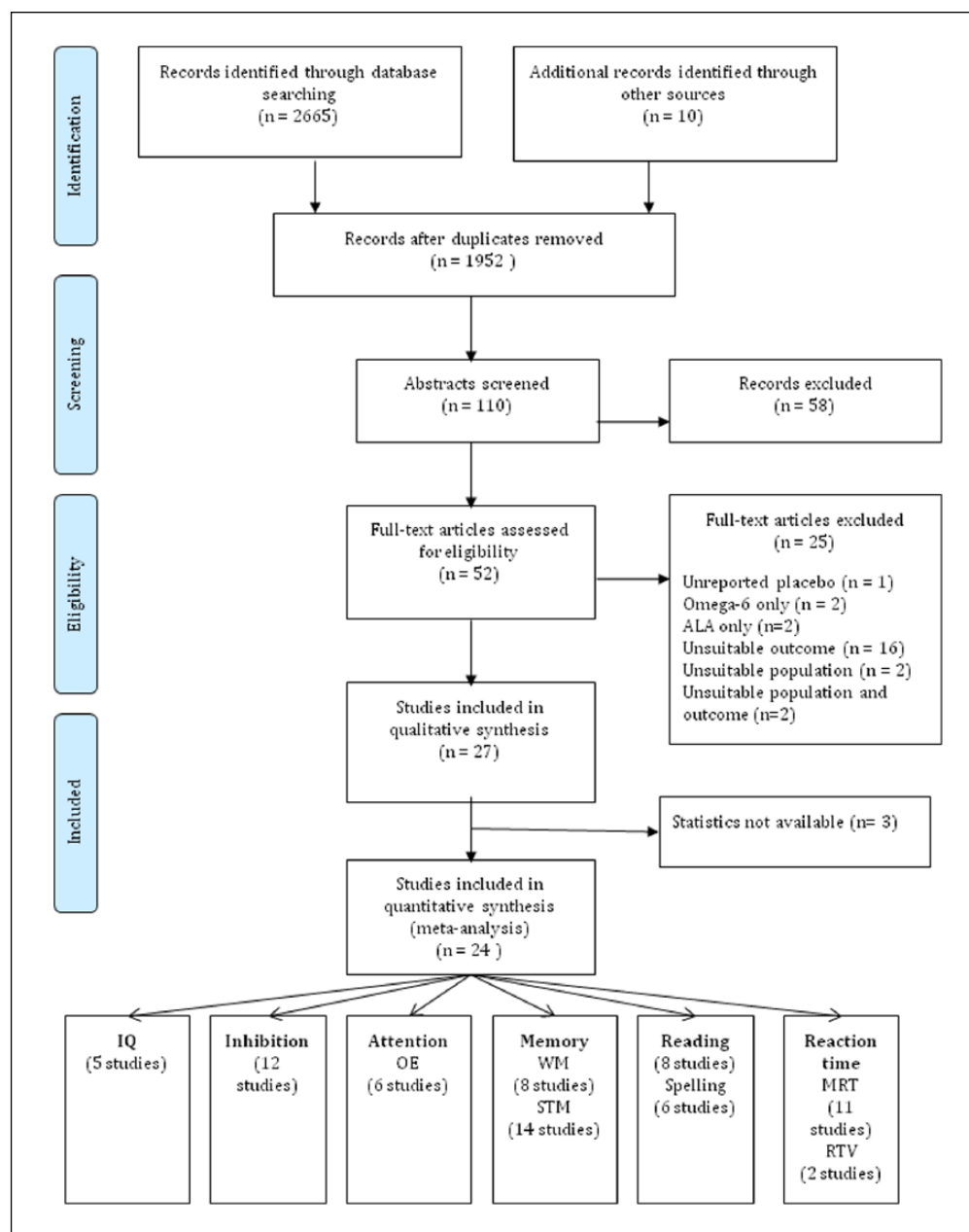
contacted authors. Missing data that remained unavailable was not imputed.

### Cognitive performance measures

Nine domains of cognitive performance, previously found to be impaired in ADHD and related disorders (Doyle et al., 2005; Frazier et al., 2004; Kuntsi et al., 2009; Willcutt et al., 2010) were measured in these studies and included in this meta-analysis (see Figure 1). Examples of the main measures and tasks were as follows: IQ measured using the Wechsler Intelligence Scale for Children (WISC; Wechsler, 1991); commission errors (the inability to withhold a pre-potent response) on computerised tasks for inhibition (e.g. continuous performance tasks); omission errors (failing to respond when a response is required) on computerised attention tasks (e.g. test of variables of attention (TOVA) (Greenberg and Kindschi, 1996)) for attention (omission errors); digit span backwards (recalling a string of numbers backwards) for working memory; immediate or delayed word recall for short term memory; reading and spelling using subtests of the Wide Range Achievement Test (WRAT; Wilkinson and Roberts, 2006); mean reaction time (speed of responding) and reaction time variability (the variability in the speed of responding) during attention tasks (e.g. TOVA) (see Supplementary Material, Table S1 for a detailed list of cognitive measures).

### Statistical analyses

Analyses were carried out in STATA (StataCorp, 2009) on the whole sample, the TD and ADHD+RD subgroups separately (with a further analysis of adults and children separately in the TD group) and then for the secondary subgroup analysis (see subgroup analysis section). Where a study contained two active groups which were both eligible for inclusion (for example when the active groups differed in the dose of *n*-3 PUFA), they were combined (with the method presented in the *Cochrane Handbook*: section 16.5.4; Higgins and Green, 2011). Effect sizes were estimated as the standardised mean difference (SMD); calculated as the mean pre-to post-treatment change, minus the mean pre-to post-placebo group change, divided by the pooled pre-test standard deviation (SD) with a bias adjustment (Morris, 2007). Effect sizes were classified according to Cohen's *d* (0.2=small, 0.5=medium, 0.8=large; Cohen, 1988). Where SD was not provided, it was calculated from sample size, *p*-values, *t*-values, standard error (SE) or 95% confidence intervals (CIs). For individual studies that contributed multiple assessments for one cognitive domain, a single SMD was derived from a meta-analysis of these assessments (see Supplementary Material, Table S1) hence an individual study was counted only once per cognitive domain. Cross-over trials were treated as parallel group trials using the pre-cross-over data, because insufficient data were provided to permit analysis of within-individual change (e.g. no correlations of scores between conditions). This approach is considered conservative (studies are under-rather than over-weighted) and is equivalent to setting the between-condition correlation to zero (Elbourne et al., 2002). SMDs in each domain were combined using the inverse-variance method where the reciprocal of their variance is used to weight the SMD



**Figure 1.** Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) flow diagram.

ALA: alpha-linoleic acid; IQ: intelligence quotient; MRT: mean reaction time; OE: omission error; RTV: reaction time variability; STM: short term memory; WM: working memory.

from each trial before being combined to give an overall estimate (Higgins and Green, 2011). Given the between-study heterogeneity in terms of study design, participant characteristics and outcome measures, we chose a priori to use random effects models (Field and Gillett, 2010). When setting the significance level, we corrected for nine domains of cognition (Bonferroni correction set at  $p < 0.006$ ) despite the fact that the primary analysis was performed in the total sample and also separately for the ADHD+RD and TD groups (i.e. more than nine statistical tests were conducted), because the cognitive tests are highly correlated. The above  $p$ -value (0.006) was considered indicative and not evidence of association for the post-hoc analyses. A nominal

level of significance was set at  $p < 0.05$ . The  $I^2$  statistic assessed heterogeneity between studies. Publication bias was assessed using the Egger regression asymmetry test (and inspection of the regression asymmetry plot) and the Begg adjusted rank correlation test. Meta-regression was used to examine the association between treatment effect and (a) trial duration and (b) dose of EPA and DHA. Four studies contained two active groups (Jackson et al., 2012; Kennedy et al., 2009; McNamara et al., 2010; Milte et al., 2012), therefore the average dose of EPA and DHA was taken across the two groups for the meta-regression and for the 'adequate EPA' subgroup analysis (see 'subgroup analyses' section).

## Subgroup analyses

1. Strict inclusion: all studies that met our inclusion criteria were included in the primary analysis (as above). As two studies used supplementation with carnosine (Kairaluoma et al., 2009) or vitamins (Kirby et al., 2010) in addition to *n*-3 PUFA we performed subgroup analyses excluding these two studies.
2. PUFA deficient: it is proposed that only participants who are deficient in *n*-3 PUFA will benefit from treatment. The analysis was therefore run in four studies that supplemented: children of low socio-economic status who had low fish intake (defined in the paper as 'virtually no intake of fatty fish and a very low intake of lean fish,' Dalton et al., 2009, section 2.1), adults with low *n*-3 PUFA intake (less than ~200 mg EPA+DHA/wk, Stonehouse et al., 2013), malnourished children (53% consumed <1 portion fish a week, 39% one portion a week and 8% ≥2 portions a week, Portillo-Reyes et al., 2014) and children with ADHD who were deficient in *n*-3 PUFA (participants were selected with thirst/skin problems indicative of *n*-3 PUFA deficiency, blood analysis showed these participants to have significantly lower *n*-3 PUFA compared with a TD control group, Stevens et al., 2003).
3. High quality: quality appraisal demonstrated the majority of studies to have design errors therefore the analysis was re-run in the eight studies whose overall risk score was low (and were therefore deemed high quality) (Supplementary Material, Table S3 and S4) (Jackson et al., 2012; Kairaluoma et al., 2009; Karr et al., 2012; Kennedy et al., 2009; Richardson and Montgomery, 2005; Richardson et al., 2012; Stonehouse et al., 2013; Vaisman et al., 2008).
4. Cognitive impairment: heterogeneity in cognitive impairments across study populations may reduce the effect size of treatment response. The analysis was run in four studies which included those with more homogeneous cognitive deficits. Milte et al. (2012) tested children with ADHD whose literary performance was behind their year level at school. Vaisman et al. (2008) tested children with ADHD who also performed poorly on a continuous performance test. Richardson et al. (2012) tested a sub-group of the poorest readers (<20<sup>th</sup> centile from the total sample) and Kairaluoma et al. (2009) tested children with dyslexia.
5. Adequate EPA: a significant association between dose of EPA (but not DHA) and improvement in ADHD symptoms has previously been found (Bloch and Qawasmi, 2011). Given this, it has been suggested that EPA may be more active than DHA in terms of its effect on brain and behaviour. The analysis was therefore run in the 14 studies which supplemented participants with >100 mg EPA (this cut-off was estimated from Figure 3 in Bloch and Qawasmi's paper) (Antypa et al., 2009; Gustafsson et al., 2010; Hamazaki et al., 1996; Jackson et al., 2012; Kairaluoma et al., 2009; Karr et al., 2012; Milte et al., 2012; Parletta et al., 2013; Portillo-Reyes et al., 2014; Richardson and Montgomery, 2005; Sinn et al., 2008; Stonehouse et al., 2013; Vaisman et al., 2008; Widenhorn-Müller et al., 2014).

## Results

### Selection of studies

The search strategy (conducted by REC) identified 1952 publications. Of these, 110 relevant abstracts were screened, of which 58 were excluded because the studies were not an RCT ( $n=34$ ), or they used an unsuitable outcome (e.g. looked only at treatment effects on PUFA blood levels) ( $n=10$ ), population ( $n=12$ ), supplement ( $n=1$ ), or study design ( $n=1$ ). Fifty-two full text articles were subsequently quality appraised (by REC and CT) and 25 excluded because of failure to report the placebo group ( $n=1$ ), supplementation with omega-6 ( $n=2$ ) or ALA ( $n=2$ ) only, use of unsuitable outcome measures ( $n=16$ ) (e.g. only measured behavioural outcomes), unsuitable population ( $n=2$ ) or unsuitable outcome and population ( $n=2$ ) (Supplementary Material, Table S5 lists the excluded studies). Of the 27 trials suitable for inclusion, after writing to the authors of studies with missing data, the statistical information required for meta-analysis was available for 24 studies, which made up the final dataset used in the meta-analysis (Figure 1 and Supplementary Material, Table S6).

### Quality and characteristics of studies included in qualitative synthesis

Randomisation was explicitly described in 20 studies and allocation concealment in 17 studies. In the remainder this was absent or unclear. All studies were double blind apart from one, where the chief investigator was unblinded (although did not collect/analyse data) (Dalton et al., 2009). Inadequate allocation concealment in two studies meant participants were aware they were in different groups (Baumgartner et al., 2012; Dalton et al., 2009). Above-chance guessing (70%) of group allocation occurred in another study (Milde et al., 2012). Drop-outs ( $n=5/25$ ) occurred only in the placebo group in one study (Portillo-Reyes et al., 2014). Reasons for drop-outs were not given in three studies (Benton et al., 2013; Kirby et al., 2010; Ryan and Nelson, 2008) despite one having more than double the amount of drop-outs in the active group (Ryan and Nelson, 2008). In one study the distribution of drop-outs between the placebo and active groups was not given (Gustafsson et al., 2010). In one study the *n*-3 PUFA supplements were taken only four days per week (Baumgartner et al., 2012). In two studies the active groups took supplementation with carnosine or vitamins in addition to *n*-3 PUFA (Kairaluoma et al., 2009; Kirby et al., 2010) (Supplementary Material, Table S3 and S4). Study characteristics are detailed in Supplementary Material, Table S7.

### Quantitative meta-analysis

Of the 27 studies included in the qualitative synthesis, pre- and post-treatment means and SDs were not available for three studies (Hirayama et al., 2004; Long and Benton, 2013; Ryan and Nelson, 2008) therefore 24 studies were included in the meta-analysis. Omega-3 PUFA supplementation had no significant effect on any of the nine domains of cognitive performance in either the whole sample or the ADHD+RD or TD group (when analysed as a whole and by adults and children) separately. An effect on working memory in the ADHD+RD group approached significance (three studies,  $n=506$ ) (SMD=0.23; 95% CI: -0.001–0.46,  $z=1.95$ ,  $p=0.05$ ) with no heterogeneity ( $\chi^2=3.03$ ,  $I^2=33.9\%$ ,



Table 1. Main effects of meta-analysis for the whole sample (combined typically developing (TD), attention deficit hyperactivity disorder (ADHD)+related disorder (RD) group).

Domain	<i>n</i> studies	<i>n</i> participants	SMD	95% CI	Heterogeneity	
					<i>p</i>	<i>I</i> <sup>2</sup> (%)
IQ	5	434	0.14	−0.07–0.35	0.28	20.9
Inhibition	12	809	−0.04	−0.22–0.14	0.08	38.7
Attention (omission errors)	6	321	−0.13	−0.33–0.07	0.96	0.0
Memory (working memory)	8	1308	0.09	−0.01–0.18	0.40	3.9
Memory (short-term memory)	14	1914	0.07	−0.01–0.15	0.15	29.0
Reading	8	1579	0.02	−0.06–0.09	0.62	0.0
Spelling	6	1167	0.03	−0.09–0.15	0.39	5.0
Reaction time (mean reaction time)	11	1035	−0.002	−0.12–0.12	0.33	12.5
Reaction time (reaction time variability)	2	91	0.29	−0.70–1.28	0.02 <sup>a</sup>	82.0

<sup>a</sup>Significant at  $p < 0.05$ .

$p = 0.22$ ). Main effects from the meta-analysis are summarised in Table 1 and Figure 2 and a detailed description of results is available in Supplementary Material, Section S1.

In the subgroup of those who were *n*-3 PUFA deficient a small treatment effect was found for short-term memory (three studies,  $n = 331$ , SMD = 0.26; 95% CI: 0.09–0.43,  $z = 3.02$ ,  $p = 0.003$ ) with no heterogeneity ( $\chi^2 = 2.67$ ,  $I^2 = 25.1\%$ ,  $p = 0.26$ ). In those who met strict inclusion criteria, a small treatment effect emerged for working memory after exclusion of one study which supplemented with vitamins (Kirby et al., 2010) (seven studies,  $n = 960$ , SMD = 0.15; 95% CI: 0.03–0.27,  $z = 2.48$ ,  $p = 0.01$ ) with no heterogeneity ( $\chi^2 = 4.18$ ,  $I^2 = 0.0\%$ ,  $p = 0.65$ ). In studies that supplemented with adequate EPA a small treatment effect emerged for working memory (five studies,  $n = 510$ , SMD = 0.19; 95% CI: 0.04–0.34,  $z = 2.44$ ,  $p = 0.02$ ) with no heterogeneity ( $\chi^2 = 3.52$ ,  $I^2 = 0.0\%$ ,  $p = 0.47$ ). Although the latter two treatment effects did not withstand correction for multiple testing (adjusted  $p < 0.006$ ). No other significant effects were found. Supplementary Material, Table S8 details the results of the subgroup analysis.

Significant heterogeneity was present in the TD group for IQ ( $\chi^2 = 4.12$ ,  $I^2 = 75.8\%$ ,  $p = 0.04$ ) and short-term memory ( $\chi^2 = 17.66$ ,  $I^2 = 49.0\%$ ,  $p = 0.04$ ). In the TD-child sample for inhibition ( $\chi^2 = 3.92$ ,  $I^2 = 74.5\%$ ,  $p = 0.05$ ), in the TD-adult sample for mean reaction time ( $\chi^2 = 8.34$ ,  $I^2 = 64.0\%$ ,  $p = 0.04$ ) and in ADHD+RD participants for reaction time variability ( $\chi^2 = 5.54$ ,  $I^2 = 82.0\%$ ,  $p = 0.02$ ). In the sub-group analyses heterogeneity was found in those who were PUFA deficient for inhibition ( $\chi^2 = 8.16$ ,  $I^2 = 75.5\%$ ,  $p = 0.02$ ) and mean reaction time ( $\chi^2 = 5.00$ ,  $I^2 = 79.3\%$ ,  $p = 0.03$ ). Meta-regression found no effect of trial duration or dose of EPA or DHA on any of the eight domains of cognitive performance (there were not enough studies to examine this for reaction time variability (RTV)).

There was evidence of publication bias in working memory (Egger test only) ( $\beta = 1.87$ , SE = 0.58,  $t = 3.23$ ,  $p = 0.02$ , 95% CI: 0.45–3.29) but not in any of the other eight domains of cognitive performance.

## Discussion

This systematic review and meta-analyses examined the efficacy of *n*-3 PUFA supplementation on cognitive performance

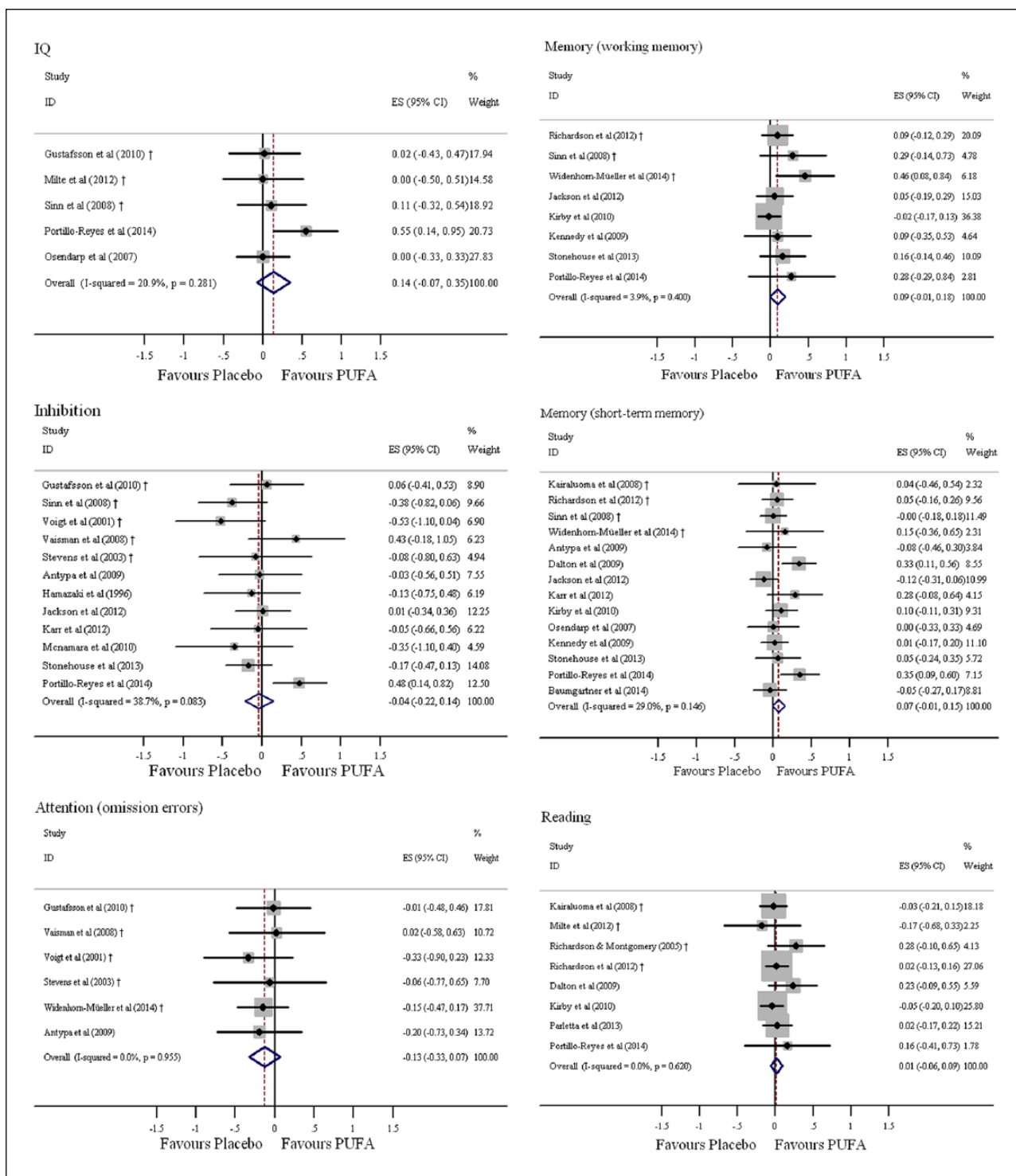
measures in school aged children and adults who were typically developing (TD) or had ADHD or a related neurodevelopmental disorder (ADHD+RD). We did not find an effect of *n*-3 PUFA supplementation on cognition in either the whole sample or the TD (analysed as a whole and by adults and children separately) or the ADHD+RD group when analysed separately. In the subgroup analyses a small treatment effect emerged for short-term memory in those with low *n*-3 PUFA and for working memory, after removal of a study which supplemented with vitamins (Kirby et al., 2010) and in those studies that supplemented with adequate EPA. However, both the effects on working memory were only nominally significant and were driven by the outcome of one cognitive measure in a small sample of 61 children with ADHD (Widenhorn-Müller et al., 2014).

There was no evidence of heterogeneity in the whole sample. Nominally significant heterogeneity was found in a number of sub-analyses (TD, ADHD+RD, TD-adult, TD-child, PUFA-deficient and high quality studies) this is most likely due to the smaller number of studies included in these analyses. Meta-regression found no effect of trial duration or EPA or DHA dose across any of the eight domains of cognitive performance (there were not enough studies to examine this for RTV). Evidence of publication bias was found only for working memory. We conclude on the basis of these data that there is no evidence of an effect of *n*-3 PUFA supplementation on cognitive performance in typically developing individuals or those with ADHD and related disorders. There is marginal evidence of benefit in those who are *n*-3 PUFA deficient. Evidence for those that met strict inclusion criteria or that supplemented with adequate EPA was much weaker.

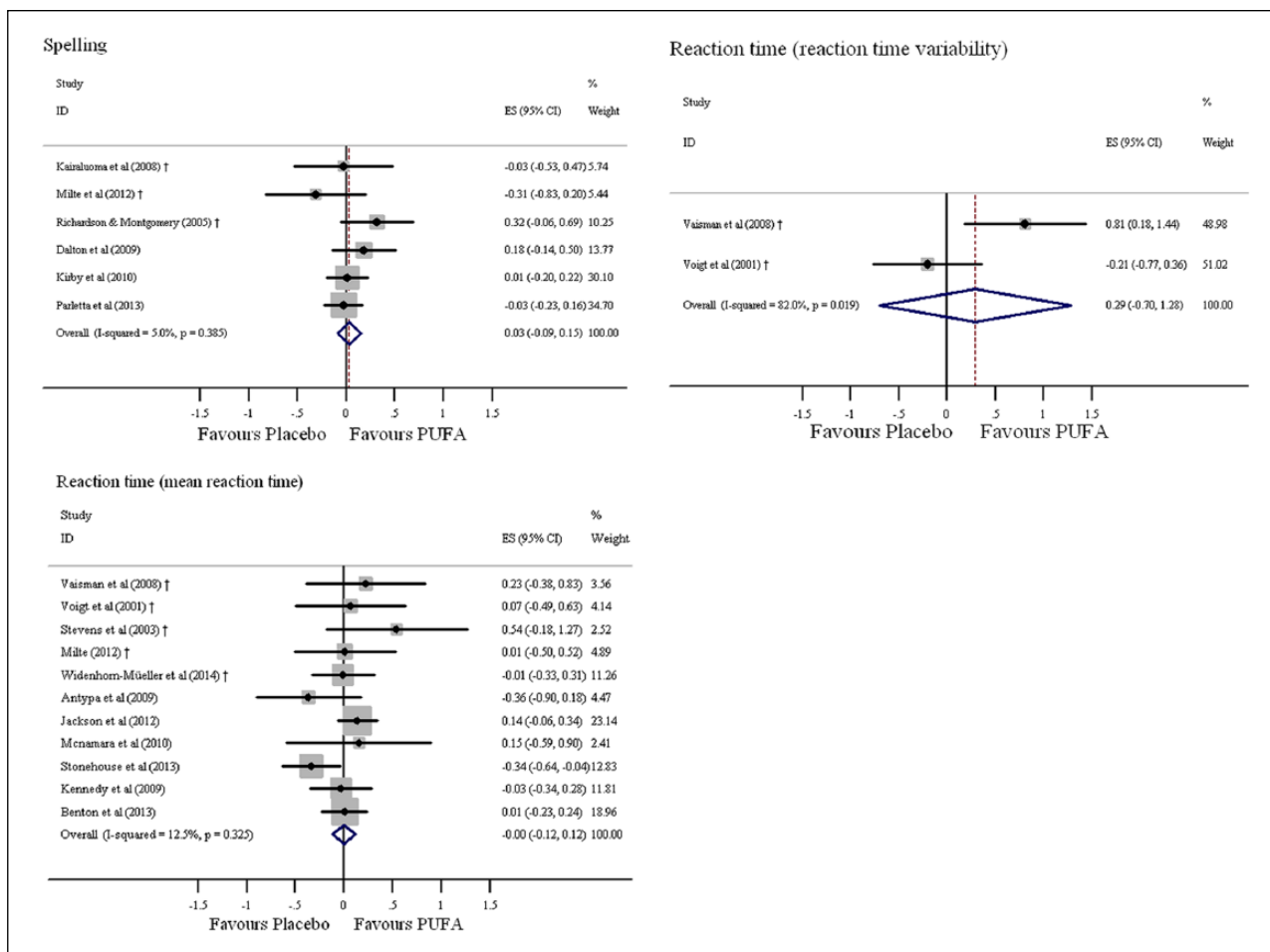
A small improvement (which withstood correction for multiple testing) in short-term memory was found across four studies (in TD and ADHD+RD populations) which supplemented those with low *n*-3 PUFA (Dalton et al., 2009; Portillo-Reyes et al., 2014; Stevens et al., 2003; Stonehouse et al., 2013). Results from one study in malnourished children also found improvements in IQ following supplementation (Portillo-Reyes et al., 2014). This is in line with the suggestion that treatment effects on cognitive performance may occur only in those with low *n*-3 PUFA levels at baseline. However only four studies could be included in this subgroup (Dalton et al., 2009; Portillo-Reyes

et al., 2014; Stevens et al., 2003; Stonehouse et al., 2013) and whilst three of them measured PUFA-blood levels, only one of these examined blood-PUFA deficiency. Stevens et al., (2003) found reduced *n*-3 PUFA status in the ADHD study participants compared to TD controls. Therefore it cannot be certain that the

other study participants were *n*-3 PUFA deficient. The subgroup analysis on those with low *n*-3 PUFA status found treatment effects in only one of the five cognitive performance domains. Therefore, although promising, further trials are needed before drawing any firm conclusions.



(Continued)



**Figure 2.** Forest plots for the meta-analyses in the whole sample across the nine domains of cognition. Studies without † indicates typically developing (TD) group. † indicates attention deficit hyperactivity disorder (ADHD)+ related disorder (RD) group. CI: confidence interval; ES: effect size; ID: identification; IQ: intelligence quotient; PUFA: polyunsaturated fatty acid.

There was no evidence of an effect of *n-3* PUFA supplementation in TD individuals and those with ADHD+RD. This is in line with the inconsistent findings from individual studies, with few positive findings remaining significant after correction for multiple testing (Antypa et al., 2009; Dalton et al., 2009; Hirayama et al., 2004; Jackson et al., 2012; Parletta et al., 2013; Sinn et al., 2008; Vaisman et al., 2008; Voigt et al., 2001). It is also in line with findings from the three studies included in the qualitative but not quantitative synthesis. Hirayama et al., (2004) found no effect on memory, attention or inhibition after eight weeks of supplementation in children with ADHD. Ryan and Nelson, (2008) found no effect on attention or inhibition after four months supplementation in healthy children. Long and Benton, (2013) found no effect on inhibition after three months supplementation in healthy adult males. This conclusion goes against previous narrative reviews which have suggested *n-3* PUFA supplementation to improve cognitive performance (Assisi et al., 2006; Bryan et al., 2004; Horrocks and Yeo, 1999; Stonehouse et al., 2013). However, while these reviews highlighted interesting findings, they failed to provide a critical analysis in light of the mixed results on performance measures.

There are several important limitations to be considered before drawing conclusions. This study was limited by substantial between study variation with respect to patient groups, assessment procedures, outcome measures, treatment formulations, and quality in methods adopted for the different studies, necessitating the use of random effects models that produced wider confidence intervals. Due to reporting deficiencies the present study used pre-treatment SD instead of SD of the change (the difference before and after the intervention) in the calculation of effect size (Morris, 2007). This could have resulted in an underestimation of the true effect size (Ortego and Botella, 2010), although a sensitivity analysis of four studies of short-term memory which gave the SD of the change gave a similar, non-significant result (see Supplementary Material, Section S1).

In accord with our predominately negative findings, it has previously been suggested that treatment for ADHD may be more effective for the behavioural symptoms of inattention and hyperactivity-impulsivity, than cognitive performance measures (Coghill et al., 2014). In line with this, previous meta analyses have found a small but significant effect of *n-3* PUFA supplementation on reducing ADHD symptoms in 699 (SMD=0.31,  $p < 0.0001$ ) (Bloch and Qawasmi, 2011) and 827 (SMD=0.21,

$p=0.007$ ) (Sonuga-Barke et al., 2013) children with ADHD. Furthermore, meta-analyses and systematic reviews have found a smaller treatment effect of stimulant medication on cognitive performance ( $\sim 0.2$ – $0.6$ ) (Coghill et al., 2014) than on ADHD symptoms ( $\sim 0.8$ – $1.0$ ) (Banaschewski et al., 2006; Faraone and Buitelaar, 2010). Several recent studies investigating the clinical response to methylphenidate found a dissociation of the treatment effects on ADHD symptoms and cognitive performance in children and adolescents with ADHD (Bédard et al., 2015; Coghill et al., 2007; Schulz et al., 2014). It is therefore suggested that different mechanisms are responsible for change in cognitive performance and change in behavioural symptoms (Coghill et al., 2007).

The lack of significant effects in the ADHD+RD group may reflect neuropsychological heterogeneity leading to a reduced effect size for individual domains of cognitive impairments, in comparison to ADHD symptoms where there is a more uniform deficit (Coghill et al., 2007; Nigg et al., 2005; Sonuga-Barke et al., 2010). For example, Vaisman et al., (2008) included children with a clinical ADHD diagnosis who also performed poorly on a continuous performance test and found a greater number of significant treatment effects on cognitive measures than studies that included those with a clinical ADHD diagnosis regardless of the baseline level of cognitive impairment (for example see Stevens et al., 2003). Although subgroup analyses across four studies which included those with more homogenous cognitive deficits (Kairaluoma et al., 2009; Milte et al., 2012; Richardson et al., 2012; Vaisman et al., 2008) failed to find treatment effects. However, given this small number of studies, further work would be required to test this specific sub-group hypotheses.

The studies used in this meta-analysis varied in supplement composition and dosage according to a previous meta-analysis higher EPA rather than DHA concentrations are associated with symptom reduction in children diagnosed with ADHD (Bloch and Qawasmi, 2011). However a subgroup analysis of those that supplemented with adequate ( $>100$  mg) EPA and a meta-regression examining the relationship between EPA dose and cognitive task performance did little to support this. We found only one small (nominally significant) treatment effect for working memory which was driven by one study (Widenhorn-Müller et al., 2014) and no effect of EPA dose on cognitive performance.

There are inherent problems with blinding in studies which supplement with  $n-3$  PUFA due to the fishy flavour of the capsules. The majority of studies did not assess blinding however above chance guessing occurred in one study that examined this (Milde et al., 2012). Identical flavouring of the placebo and active capsules must be used to reduce this limitation and the possibility of inflated effect sizes. A large number ( $n=7$ ) of the studies included in the qualitative synthesis used an olive oil placebo. Olive oil contains a high concentration of oleic acid, a precursor of oleamide that has been shown to have psychoactive properties (Richardson, 2006). Stevens et al. (2003) found their olive oil placebo to be 'active' in that the supplement did not maintain the baseline PUFA composition. An inert substance such as liquid paraffin oil could be more suitable (Peet and Horrobin, 2002).

The majority of studies used in this meta-analysis were underpowered. The treatment effect that withstood correction for multiple testing (short-term memory in those who were  $n-3$  PUFA deficient ( $SMD=0.26$ )) was small. With this modest effect size of around 0.3 we would require a sample size of around 352 participants ( $\beta=80\%$ , two-tailed  $\alpha=0.05$ ) at a nominal level of significance

and around 596 participants after correction for multiple testing ( $\beta=80\%$ , two-tailed  $\alpha=0.006$ ). In the ADHD+RD group trials ranged from 40–362 participants with only three trials above 100. Although the largest trial (Richardson et al., 2012) in healthy children underperforming in reading found treatment effects on reading in only a subgroup of those who were the poorest readers and no effect on working memory. The largest trial in children with ADHD ( $n=110$ ) (Widenhorn-Müller et al., 2014) again found only marginal evidence of a treatment effect with improvement in working memory but not in six other cognitive performance measures. Future studies should be adequately powered to detect small effects in order to clarify the presence of treatment effects.

We included only school-aged children and adults in our analysis (no trials in adolescent populations were located). The current results are therefore not generalisable to infants, adolescents or the elderly. Research has suggested similar negative results in these groups. A recent meta-analysis examined the effect of  $n-3$  PUFA on cognitive performance in healthy elderly adults and those with cognitive decline. Across 10 domains of cognitive performance, treatment effects were found for those with cognitive decline in three domains (immediate and delayed recall, attention/processing speed). However significance was only at a nominal level ( $p=0.02$ – $0.04$ ) and became non-significant after correction for multiple testing (Mazereeuw et al., 2012). A Cochrane review and meta-analysis concluded RCTs in infants to have provided little evidence for the effect of  $n-3$  PUFA on neurodevelopmental outcomes (including cognition) and inconsistent effects on visual acuity (Simmer et al., 2011).

Sex dimorphism may also be present in response to PUFA supplementation, thus analysis of samples as a whole and not by sex could potentially mask effects. One study found improvement in episodic memory in women and working memory in men (Stonehouse et al., 2013), potentially reflecting gender differences in problem-solving strategies. However, these findings were not corrected for multiple testing and further evidence would be required to examine the question of sex dimorphism in the cognitive response to  $n-3$  PUFA supplementation.

Length of supplementation has also been proposed as a factor. In the current study only three trials were of six months or longer (Dalton et al., 2009; Osendarp et al., 2007; Stonehouse et al., 2013). Across two of these studies treatment effects were found on verbal learning ability, memory and reaction time (Dalton et al., 2009; Stonehouse et al., 2013). Although Stonehouse et al. (2013) tested a large number of cognitive domains, the majority of which were non-significant and failed to correct for multiple testing. The longest study (12 months) failed to find any treatment effects (albeit the dosage of  $n-3$  PUFA was relatively small; Osendarp et al., 2007). The current study found no relationship between length of supplementation and effects on cognitive performance which is in line with a recent meta-analysis that found no relationship between trial duration and efficacy of  $n-3$  PUFA supplementation in reducing ADHD symptoms (Bloch and Qawasmi, 2011). This evidence suggests that outcomes may have been uninfluenced by duration.

A number of outcome measures such as accuracy on cognitive tasks could not be included in this analysis due to the measures being too diverse to combine. However results from such measures were largely negative. For example treatment effects were not found in TD children (Kennedy et al., 2009) and adults (Jackson et al., 2012) for accuracy on reaction time tasks or in



children with ADHD for speed of information processing tasks (Widenhorn-Müller et al., 2014). Although one study (Sinn et al., 2008) found a significant benefit of treatment for accuracy on a sustained attention task, overall these results are in line with current negative findings.

In conclusion we have found no evidence of an effect of *n*-3 PUFA supplementation on cognitive performance in the general population or in those with ADHD and related disorders. There was suggestive evidence of improvements in those with low *n*-3 PUFA status. In order to provide a more conclusive picture future trials should employ larger sample sizes and should focus on supplementation of those who are *n*-3 PUFA deficient. It is suggested that regulators and producers of omega-3 products should consider this evidence when promoting their products.

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